

Research Paper

A Novel Risk prediction Model for Patients with Combined Hepatocellular-Cholangiocarcinoma

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Abstract

Backgrounds: Regarding the difficulty of CHC diagnosis and potential adverse outcomes or misuse of clinical therapies, an increasing number of patients have undergone liver transplantation, transcatheter arterial chemoembolization (TACE) or other treatments.

Objective: To construct a convenient and reliable risk prediction model for identifying high-risk individuals with combined hepatocellular-cholangiocarcinoma (CHC).

Methods: 3369 patients who underwent surgical resection for liver cancer at Zhongshan Hospital were enrolled in this study. The epidemiological and clinical characteristics of the patients were collected at the time of tumor diagnosis. Variables ($P < 0.25$ in the univariate analyses) were evaluated using backward stepwise method. A receiver operating characteristic (ROC) curve was used to assess model discrimination. Calibration was performed using the Hosmer-Lemeshow test and a calibration curve. Internal validation was performed using a bootstrapping approach.

Results: Among the entire study population, 250 patients (7.42%) were pathologically defined with CHC. Age, HBcAb, red blood cells (RBC), blood urea nitrogen (BUN), AFP, CEA and portal vein tumor thrombus (PVTT) were included in the final risk prediction model (area under the curve, 0.69; 95% confidence interval, 0.51-0.77). Bootstrapping validation presented negligible optimism. When the risk threshold of the prediction model was set at 20%, 2.73% of the patients diagnosed with liver cancer would be diagnosed definitely, which could identify CHC patients with 12.40% sensitivity, 98.04% specificity, and a positive predictive value of 33.70%.

Conclusions: Herein, the study established a risk prediction model which incorporates the clinical risk predictors and CT/MRI-presented PVTT status that could be adopted to facilitate the diagnosis of CHC patients preoperatively.

Key words: combined hepatocellular-cholangiocarcinoma, risk prediction model, liver cancer, preoperation.

Introduction

Combined hepatocellular-cholangiocarcinoma (CHC) is a rare primary liver malignancy with dual histologic differentiation: hepatocellular and biliary

epithelial features[1, 2], accounting for 0.4-14.2% of all primary liver cancers in Asia and Western countries[1-4]. Since Allen and Lisa firstly reported

this type of liver cancer, two histopathological classification schemes, Allen-Lisa and Goodman classifications, have been used for CHC classification [3, 5]. Over the past two decades, a growing number of retrospective studies have described this rare tumor [6-8]. To date, most studies primarily focused on the demographic and clinical characteristics of CHC. However, in contrast to the established preoperative diagnostic prediction model of hepatocellular carcinoma (HCC) [9, 10], little is known about the preoperative diagnostic prediction of CHC.

The clinical features of CHC are similar to either HCC or intrahepatic cholangiocarcinoma (ICC), which implies that the risk factors (tumor biomarkers, etiology, characteristics and imaging tests) of CHC overlapped with those of HCC or ICC [1, 6, 11-14]. But the imaging features of CHC patients didn't present typical features compared with HCC or ICC patients, which manifested that preoperative diagnosis of CHC is very difficult [15]. Thus, patients who were diagnosed with CHC by pathological tests may potentially receive liver resection, transcatheter arterial chemoembolization (TACE) or liver transplantation according to HCC guideline [16-18], as TACE was less effective in the treatment of CHC patients than radical resection [1] [7]. In addition, CHC was associated with poorer prognosis and higher rate of tumor recurrence after liver transplantation [19]. Considering the poor response to TACE, the scarcity of liver donation [20] and the difficulty of preoperative diagnosis, it is necessary to establish a risk prediction tool that identifies CHC patients and available for clinical practice.

In the present study, patients were categorized into CHC and non-CHC cohorts. The aim of the present study was to investigate the essential differences between CHC and non-CHC populations based on epidemiologic and clinical characteristics at the time of tumor diagnosis. We also developed and validated the clinical preoperative CHC risk prediction model using bootstrap method.

Material and methods

Patients

Between March 1993 and December 2014, 4245 patients with liver cancer were surgically treated and screened for enrollment in the Department of Liver Surgery, Zhongshan Hospital. The standard technique was adopted for hepatic resection [20], and the pathology of liver cancer was histologically defined according to World Health Organization criteria [21, 22]. Preoperative imaging studies included chest X-ray, abdominal ultrasonography, contrast-enhanced computer tomography (CT) or

magnetic resonance imaging (MRI) and, in some cases, hepatic arteriography. Extrahepatic metastasis was diagnosed using chest X-ray, chest CT, and bone scintigraphy. Liver function was assessed using serum biochemical data, ascites, and prothrombin time. Patients with Child-Pugh class A underwent major hepatic resection (more than 3 segments) [23]. The flow chart of patient exclusion was shown in **Figure 1**. In the final dataset, 250 patients were classified as a CHC cohort, and 3119 patients were classified as a non-CHC cohort. The CHC cohort presented with 82.0% (205/250) complete cases, and non-CHC cohort presented 98.1% (3059/3119). This study was approved by the institutional review board of Zhongshan Hospital and complied with the standards of the Declaration of Helsinki and current ethical guidelines.

Follow-up

In our department, the entire patients underwent monthly follow-up in the first 6 months and every 3 months thereafter until death or dropout. Abdominal ultrasound, liver function tests, serum alpha-fetoprotein (AFP), carcinoembryonic antigen (CEA) and carbohydrate antigen 19-9 (CA19-9) levels were analyzed every 3 months, and abdominal magnetic resonance imaging (MRI) scanning was performed every 6 or 12 months. Recurrence was primarily diagnosed based on imaging findings from MRI or CT scans and increased serum AFP or CA19-9 levels. Chest CT and bone scintigraphy were used to evaluate extrahepatic recurrence. Depending on the type of recurrence, and liver function reserve, patients were treated with different therapies, including RFA, repeated resection, TACE, percutaneous ethanol injection (PEI), and chemotherapy for patients with extrahepatic metastatic disease.

Laboratory test and Data collection

The serum indicators and blood cell count were measured according to standard laboratory procedures. Routine examination included serum AFP, CEA and CA19-9. In all patients, hepatitis B surface antigen (HBsAg) and antibodies to hepatitis C virus were detected using standard test systems. Demographic, pathological and clinical patient data were collected. To evaluate the potential predicting variables, the clinical and imaging variables included the patient's age, sex, etiology, laboratory tests, tumor biomarkers, tumor size, tumor number and portal vein tumor thrombus (PVTT). In the initial imaging test, the existence of PVTT was confirmed using CT or MRI. All variables were available in the practice medical record at the time of initial liver cancer diagnosis.

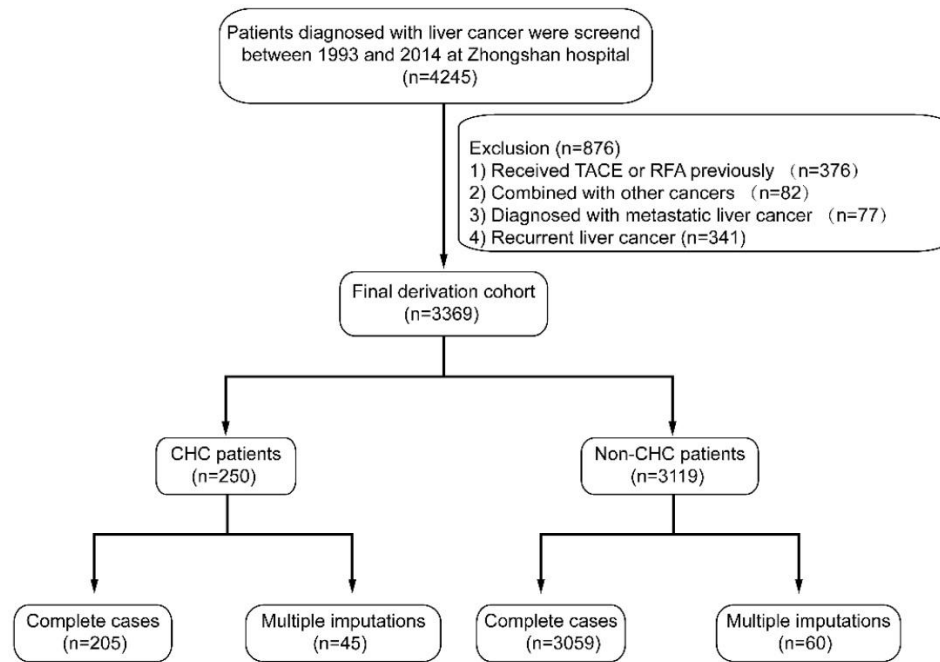


Figure 1. Flow chart of the study cohort. TACE, transcatheter arterial chemoembolization; RFA, radiofrequency ablation; CHC, combined hepatocellular-cholangiocarcinoma.

Statistical analysis

Demographic, clinical and tumor characteristics were described as summary statistics and presented as percentages or mean values. In the present study, the statistical analysis was performed according to Boursi et al [24]. Briefly, we performed multiple imputation using the Monte Carlo approach and produced 20 datasets[25-27]. The imputation of missing values was completed using Gibbs sampling[28-30]. The variables with the missing values would be predicted with other variables. Continuous variables were imputed based on the mean of the predicted values, while the categorical variables were imputed using logistic regression. For model developing procedures, univariate analysis was used for initial variable selection, and all variables with $P < 0.25$ were further evaluated using multivariable logistic regression. In each of the imputed datasets, we used a backward stepwise approach for the multivariable logistic regression with the Akaike information criterion (AIC). Predictors selected in $\geq 50\%$ of the imputation models were included in the final multivariable model, and the results of the final multivariable model were pooled in 20 imputed datasets using the mice package. For model prediction performance, discrimination and calibration were evaluated. Discrimination reflects the ability of the risk score to differentiate between patients who do and do not experience CHC. The measure is quantified by calculating the area under the receiver operating characteristic curve statistic[31]. Calibration reflects

the agreement between predicted probabilities from the model and observed outcomes. We used the Hosmer-Lemeshow test[24] to statistically determine the extent of agreement between the predicted and observed probabilities. For model validation, internal validation was adopted using a bootstrapping method[32]. The bootstrapping was performed using 100 bootstrap resamples of 3369 individuals, each time selecting variables and developing a model within the sample. The discrimination for each model was calculated both within the sample in the original cohort, enabling the calculation of the optimism according to Harrell's algorithm[33]. All analyses were performed using R software with rms, mice, and MASS packages (<http://mirrors.ustc.edu.cn/CRAN/>).

Results

Characteristics of the study cohort

The demographic and clinicopathological data for the entire cohort (3369) were shown in **Table 1**. Among the patients diagnosed with liver cancer, the mean age was 54.2 years (46.0, 62.0), men comprising 79.1% of the cohort. Based on WHO criteria, 250 (7.42%) patients were pathologically classified as CHC cohort and the remaining patients (3119 patients) were classified as the non-CHC cohort. In addition, 2489 (73.9%) patients were positive for hepatitis B surface antigen (HBsAg), and 45 (1.3%) patients were positive for the hepatitis C antibody; the mean serum AFP, CEA and CA19-9 levels were 1858.5 (3.3, 303.4), 8.1 (1.5, 3.5), 333.5 (10.9, 43.2), respectively.

Table 1. Participant characteristics and effects of CHC risk in the univariable logistic regression analysis.

Patient demographics	Entire cohort	Patients with CHC	Patients with Non-CHC	Missingness, n (%)	Crude OR (95%CI)	P value
Overall	3264	250	3119			
Host factors						
Age, yr	54.2 (46.0, 62.0)	52.0 (44.0, 60.0)	54.4 (47.0, 62.0)	1 (0.03%)	0.74 (0.56, 0.99)	0.043
Sex (male), n (%)	2665 (79.1%)	189 (75.3%)	2476 (79.4%)	1 (0.03%)	0.80 (0.60, 1.09)	0.157
Etiology						
HBsAg (positive), n (%)	2489 (73.9%)	188 (74.9%)	2301 (73.8%)	15 (0.46%)	1.08 (0.80, 1.46)	0.622
HBcAb (positive), n (%)	3013 (89.4%)	210 (83.7%)	2803 (89.9%)	18 (0.55%)	0.56 (0.39, 0.80)	0.001
HCV (positive), n (%)	45 (1.3%)	5 (2.0%)	40 (1.3%)	20 (0.61%)	1.56 (0.61, 3.99)	0.352
Liver cirrhosis (yes), n (%)	2313 (68.7%)	175 (69.7%)	2138 (68.5%)	1 (0.03%)	1.09 (0.82, 1.45)	0.544
Laboratory values						
RBC, $\times 10^{12}/L$	4.4 (4.1, 4.8)	4.3 (4.1, 4.8)	4.4 (4.1, 4.8)	14 (0.43%)	0.71 (0.58, 0.88)	0.002
HB, g/L	136.3 (127.0, 148.0)	135.3 (125.8, 149.0)	136.4 (127.0, 148.0)	12 (0.37%)	1.00 (0.99, 1.00)	0.343
WBC, $\times 10^9/L$	6.1 (4.4, 6.9)	5.9 (4.4, 6.8)	6.1 (4.4, 6.9)	37 (1.13%)	0.97 (0.92, 1.02)	0.298
PLT, $\times 10^9/L$	154.3 (105.0, 193.0)	148.9 (103.0, 186.8)	154.7 (106.0, 193.0)	12 (0.37%)	0.90 (0.66, 1.22)	0.505
AFP, ng/mL	1858.5 (3.3, 303.4)	2327.3 (5.6, 365.6)	1821.4 (3.2, 300.5)	4 (0.12%)	2.10 (1.57, 2.80)	<0.0001
CEA, $\mu g/mL$	8.1 (1.5, 3.5)	5.8 (1.2, 3.9)	8.3 (1.5, 3.5)	7 (0.21%)	1.52 (1.08, 2.14)	0.016
CA19-9, U/mL	333.5 (10.9, 43.2)	113.3 (12.1, 57.0)	350.8 (10.7, 42.7)	5 (0.15%)	1.33 (1.01, 1.75)	0.039
Bilirubin, $\mu mol/L$	18.2 (9.0, 15.7)	14.8 (9.4, 15.9)	18.5 (9.0, 15.7)	1 (0.03%)	1.04 (0.75, 1.43)	0.814
Albumin, g/L	40.7 (38.0, 43.0)	41.3 (38.0, 44.0)	40.6 (38.0, 43.0)	1 (0.03%)	1.16 (0.89, 1.52)	0.274
PA, g/L	0.3 (0.17, 0.25)	0.2 (0.17, 0.25)	0.3 (0.17, 0.25)	217 (6.65%)	0.06 (0.01, 0.69)	0.023
ALT, IU/L	51.5 (21.0, 50.0)	45.5 (20.0, 51.0)	51.9 (21.0, 50.0)	1 (0.03%)	0.97 (0.75, 1.26)	0.813
GGT, U/L	110.7 (37.0, 116.0)	102.1 (36.0, 128.3)	111.4 (37.0, 118.0)	3 (0.09%)	0.95 (0.72, 1.26)	0.736
ALP, IU/L	104.4 (66.0, 110.0)	102.5 (72.0, 115.0)	104.5 (66.0, 109.0)	13 (0.40%)	1.00 (1.00, 1.00)	0.776
APTT, s	28.8 (26.0, 31.1)	28.7 (25.5, 32.7)	28.9 (26.0, 31.1)	100 (3.06%)	0.99 (0.97, 1.02)	0.732
PT, s	12.3 (11.3, 12.7)	12.0 (11.3, 12.6)	12.3 (11.3, 12.7)	6 (0.18%)	0.93 (0.84, 1.02)	0.101
INR	1.0 (0.97, 1.08)	1.1 (0.96, 1.09)	1.0 (0.97, 1.08)	15 (0.46%)	0.79 (0.61, 1.03)	0.077
BUN, nmol/L	5.7 (4.4, 6.4)	4.8 (4.0, 5.6)	5.8 (4.5, 6.5)	63 (1.93%)	0.70 (0.63, 0.77)	<0.0001
Cr, $\mu mol/L$	72.1 (61.8, 80.6)	73.0 (61.0, 80.0)	72.1 (61.8, 80.6)	35 (1.07%)	1.00 (1.00, 1.01)	0.380
Tumor numbers (≤ 3), n (%)	2961 (87.9%)	207 (81.4%)	2754 (88.4%)	3 (0.09%)	0.99 (0.61, 1.61)	0.969
Tumor diameter, cm	6.1 (3.0, 8.0)	6.5 (3.5, 8.1)	6.1 (3.0, 8.0)	1 (0.03%)	1.03 (1.00, 1.06)	0.101
PVTT (positive), n (%)	331 (9.8%)	19 (7.6%)	312 (10.0%)	1 (0.03%)	0.70 (0.43, 1.14)	0.153

Values are presented as no. (%) or mean (Q1, Q3).

HBV, hepatitis B virus; HCV, hepatitis C virus; CH, chronic hepatitis; LC, liver cirrhosis; NL, normal liver; AFP, α -fetoprotein; CEA, carcino-embryonic antigen; CA19-9, carbohydrate 19-9; RBC, red blood cell; WBC, white blood cell; PLT, platelet; PA, pre-albumin; ALT, alanine aminotransferase; GGT, γ -glutamyl transpeptidase; ALP, alkaline phosphatase; PT, prothrombin time; INR, international normalized ratio; BUN, blood urea nitrogen; Cr, creatinine; PVTT, portal vein tumor thrombus.

Table 2. Final multivariable prediction model^a and a case example.

Predictor	β Coefficient	SE	OR	95% CI	Nmis	Fmi	Lambda	P-value
Age (<60/ ≥ 60)	-0.0268068	0.0099689	1.265	1.172-1.365	3	0.003	0.003	0.007
HBcAb (yes/no)	-0.0638200	0.0154336	0.974	0.957-0.988	NA	0.004	0.004	<0.0001
AFP (<20/ ≥ 20 ng/mL)	0.0572810	0.0093992	1.059	1.040-1.079	37	0.049	0.048	<0.0001
CEA (<5/ ≥ 5 ng/mL)	0.0354218	0.0135939	1.036	1.009-1.064	62	0.006	0.006	0.009
RBC, $\times 10^{12}/L$	-0.0276869	0.0081596	0.974	0.955-0.993	NA	0.004	0.003	<0.001
BUN, nmol/L	-0.0010271	0.0006005	0.999	0.998-1.000	NA	0.044	0.044	0.087
PVTT (yes/no)	-0.0322564	0.0151275	0.968	0.940-0.997	NA	0.001	0.0003	0.033

RBC, red blood cell; BUN, blood urea nitrogen; PVTT, portal vein tumor thrombus; OR, odds ratio. Nmis, number of missing; Fmi, fraction of missing information.

^aThe formula of the resulting logistic model is:

P probability for CHC = $e(\chi\beta) / (1 + e(\chi\beta))$.

$\chi\beta = -0.02680679 \times \text{age} - 0.06381997 \times \text{HBcAb} - 0.02768688 \times \text{RBC} - 0.00102712 \times \text{BUN} + 0.05728100 \times \text{AFP} + 0.03542176 \times \text{CEA} - 0.0322564 \times \text{PVTT} + 0.23481440$.

Case example: a 65-year-old female diagnosed with liver cancer. The tumor was less than 3cm and enhanced MRI could not distinguish the tumor from HCC or other liver cancers. The laboratory test results were HBcAb positive, AFP 207 ng/mL, CEA 75 ng/mL, RBC $5.4 \times 10^{12}/L$, and BUN 7.8 nmol/L. The imaging test of enhanced MRI suggested that the left portal vein contained tumor thrombus. According to the model, the patient's risk for CHC would be 51.2%.

In the present study, 27 candidate predictors with CHC were considered. The univariate logistic regression analysis is summarized in **Table 1**. With respect to factors associated with CHC risk, 13 variables achieved significance at $P < 0.25$, including age, sex, HBcAb, AFP, CEA, CA19-9, PA, PT, INR, BUN, tumor numbers, tumor diameter, and PVTT.

Construction of an individualized model

All factors selected above were entered into the

multivariate logistic regression analysis using a backward stepwise approach, and the results were presented in **Table 2**. Fmi represented the fraction of missing information, and lambda reflected the proportion of total variance attributable to the missing data. Predictors in $\geq 50\%$ of the 20 imputed datasets, including age (19/20, $P=0.007$), HBcAb (20/20, $P<0.0001$), AFP (20/20, $P<0.0001$), CEA (20/20, $P=0.009$), PVTT (20/20, $P=0.03$), RBC (19/20, $P<0.001$) and BUN (20/20, $P=0.08$), were selected in the final

model. The area under the curve (AUC) of the model was 0.69 (95% CI, 0.51 to 0.77, **Figure 2**) and the *P*-values for the Hosmer-Lemeshow goodness of fit test were close to 1.00 in the 20 imputed datasets.

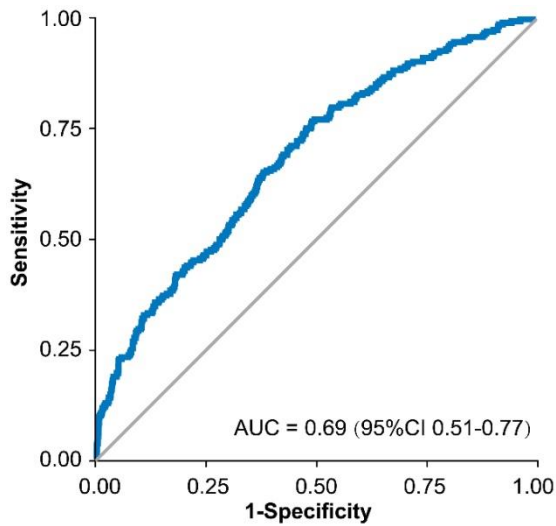


Figure 2. Receiver operating characteristic curve of the final model.

Performance and validation of the risk prediction model

To validate the performance of this risk model, we performed internal validation using the bootstrap method with 100 repetitions. The model was validated with 20 imputed datasets, and the bootstrap-corrected C-index was between 0.66 and 0.69, which demonstrated good accuracy for assessing CHC risk. The calibration curve demonstrated that the apparent probability of CHC was close to the ideal probability (**Figure 3**). The optimism of C-index ranged from 0.0072 to 0.0122.

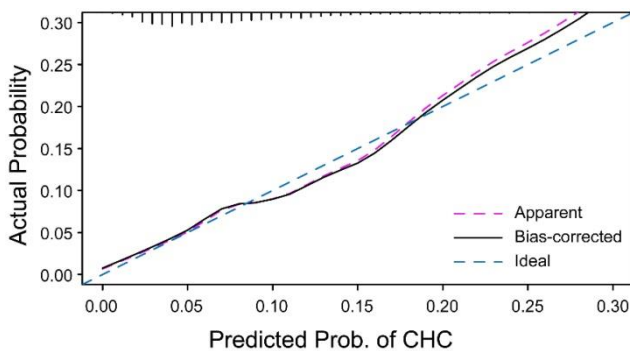


Figure 3. Calibration curve for the CHC risk model. With the dataset, the apparent calibration curve is similar to the ideal calibration curve for the CHC risk model.

Predictive curve for the CHC risk model

The predictive curve for the CHC risk prediction model was shown in **Figure 4**. If the risk threshold for

further CHC examination was set at 20%, then 2.73% of the liver cancer population would undergo screening, and the sensitivity, specificity and positive predictive value of the model would be 12.40%, 98.04%, and 33.70%, respectively. For a risk threshold of 5%, 60.7% of the liver cancer population would undergo screening, and the corresponding sensitivity, specificity and positive predictive values would be 82.00%, 40.97%, and 10.02%, respectively. The sensitivity, specificity and positive predictive value for four different probability cutoffs are presented in **Table 3** (5%, 10%, 15% and 20%).

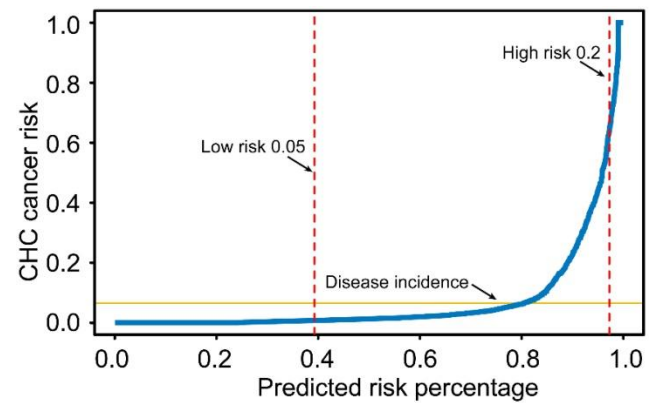


Figure 4. Predictive curve for the CHC risk model in table 3. The risk thresholds for <5% for low risk and >20% for high risk are shown.

Table 3. Model diagnostic performance at different predicted probability cut-offs.

Probability cut-off	Diagnostic performance, %
5%	
Sensitivity	82.00
Specificity	40.97
PPV	10.02
10%	
Sensitivity	45.20
Specificity	77.22
PPV	13.99
15%	
Sensitivity	23.20
Specificity	93.91
PPV	23.39
20%	
Sensitivity	12.40
Specificity	98.04
PPV	33.70

PPV=positive predictive value

Discussion

In the present study, we constructed and validated a novel risk prediction model for CHC patients. The model is based on demographic, clinical and imaging characteristics which is available at the time of tumor diagnosis. The final model presented good discrimination (0.69; 95%CI, 0.51-0.77) with

negligible optimism (0.0072 to 0.0122) and adequate goodness of fit. To our knowledge, this is the first study to establish a preoperative risk prediction model of CHC.

In clinical practice, CHC is clinically asymptomatic until the advanced stage, and these patients lose the opportunity to receive curative surgical resection [34]. Considering its dual tumor features, the existing tumor biomarkers (including CA19-9, AFP and CEA) and imaging techniques (including ultrasound, CT or MRI) hinder the accurate preoperative diagnosis of CHC. Previous studies existed two drawbacks: the first is that was small populations (12~103), and the second shortage was that these studies generally focused on the clinical and pathological characteristics of CHC[6, 7, 11, 20, 35]. These issues make it difficult to identify the potential diagnostic predictors for CHC. Other than harboring the largest data set, we included 27 potential clinical risk factors and identified 7 risk predictors using multivariate logistic regression analysis based on a backward stepwise approach. For the missing data, the multiple imputation method was performed, and 20 imputed datasets were produced. The lambda values of 7 predictors were low (**Table 2**), indicating that the effect of missing data on the final model is negligible.

Due to the difficulty of preoperative diagnosis, these patients may initially be defined as HCC or ICC, and may receive TACE, radical resection, RFA or liver transplantation based on the experience of the clinicians.[19] Although CHC patients could acquire survival benefit similar to HCC patients who underwent liver transplantation[18], the study cohort was only confined to a strict selection criteria and small populations. In addition, although TACE may be effective for prolonging survival in unresectable CHC, the survival after TACE is significantly dependent on tumor size, vascularity and the presence or absence of portal vein invasion. Thus, a novel prediction models and screening methods for detection of CHC are urgently needed.

Conducting preoperative pathological diagnosis using costly and/or invasive tests (such as biopsy) would not be a practical approach. Recently, a predictive curve has been used to evaluate the efficiency of given biological markers, to assess the fit of models and estimate the clinical utility of a model in specific population[24, 36-39]. The prediction model shown in the present study provides a low-cost, convenient and low-risk solution to this problem by identifying high-risk individuals for definitive preoperative diagnosis. For example, using a 10% predicted risk of CHC as the threshold, approximately 23.98% patients would receive a

definitive procedure. Despite the high screening percentage at 10%, a majority of patients with typical imaging or clinical characteristics for HCC or ICC would be excluded. If the threshold for proceeding with definitive testing was set as 20%, then only 2.73% of the entire cohort would receive the definitive procedure. Notably, this strategy would enable the efficient screening of these cancers.

In the present study, 7 different variables were identified as risk predictors for CHC. First, age has been demonstrated as an independent risk factor for CHC. In previous reports, the age of CHC patients was lower than that of HCC or ICC patients[7]. Age was an important variable for predicting disease-specific survival after liver resection[40], and similar results were presented in gastric[41], breast[42, 43], gallbladder[44] and colorectal cancers[45]. As a classical serological marker of HBV infection, HBcAb was identified as another risk predictor. Positive HBcAb correlated with vascular invasion and poor RFS in HCC patients after curative resection[46, 47]. In addition, these studies focused on the characteristics in CHC patients with respect to HCC or ICC patients. Elevated serum AFP and CEA were associated with poorer prognosis for ICC or HCC patients[48-51]. Although tumor biomarkers (AFP, CEA and CA19-9) may not be specific for CHC, these proteins are of diagnostic value. In the present study, AFP and CEA were demonstrated as risk predictors with an odds ratio of 1.06 and 1.04, respectively. The CHC cohort presented higher serum AFP level (2327.3, 5.6 to 365.6) and lower CEA level (5.8, 1.2 to 3.9) than that of the non-CHC cohort, while in other reports, the AFP levels in CHC patients were significantly lower compared with other groups. The result are consistent with the results of previous reports.

There are several limitations to the proposed risk-prediction model. First, the CHC cohort was derived from a single-center database in China with HBsAg positive patients accounting for 74.9% and HCV positive patients accounting for 2.0%; therefore, this model may not be applicable to other populations in western countries and further validation is warranted. Considering the small size population of CHC in previous investigations, external validation from other liver centers may not be feasible. Second, genetic polymorphisms, epigenetic changes, and information regarding circulating tumor cells, which may be essential for accurate diagnosis, were not included in our study. However, the risk-prediction model presents optimal predictive power. Finally, the model was not intended as a definitive diagnostic test but rather as part of a sequential approach to identify individuals at high-risk for CHC.

Conclusion

In conclusion, we established a risk-prediction model based on clinical, imaging characteristics, and etiology factors for predicting CHC. To date, there existed no defined criteria to discriminate patients with CHC from other liver cancer patients yet, this novel model may be useful in clinical practice as it may tailor appropriate therapy for each CHC patients.

Abbreviations

CHC: combined hepatocellular-carcinoma; TACE: transcatheter arterial chemoembolization; ROC: receiver operating characteristic curve; HCC: hepatocellular carcinoma; ICC: intrahepatic cholangiocarcinoma; CT: computer tomography; MRI: magnetic resonance imaging; PVTT: portal vein tumor thrombus; AIC: akaike information criterion.

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Ethical standards

The study with clinical data was approved by the Ethics Committee of the Zhongshan Hospital, Fudan University. We clarify that all clinical data in this study was collected in patients who had given written informed consent.

Competing Interests

The authors have declared that no competing interest exists.

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